### Mixture of Gaussian Regressions with logistic weights and conditional density estimation

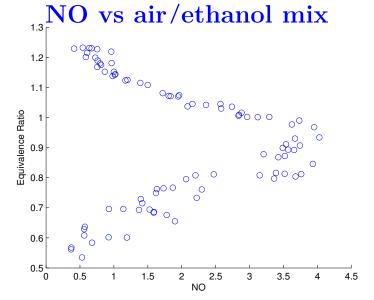
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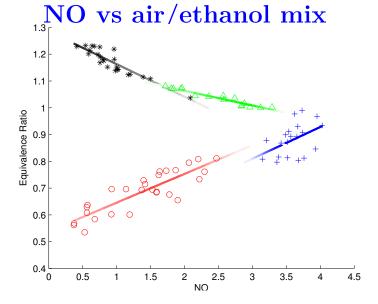
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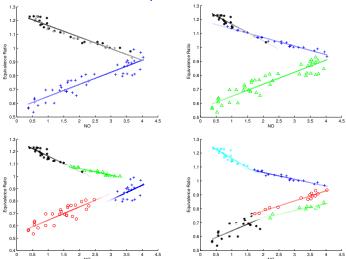


Regression



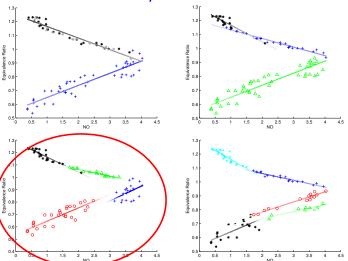
Regression Mixture

# NO vs air/ethanol mix



- Regression Mixture
- Mixtures of Gaussian regressions with logistic weights

# NO vs air/ethanol mix



- Regression Mixture
- Mixtures of Gaussian regressions with logistic weights
- Model selection

# Statistical modeling

- Data:  $(X_i, Y_i)_{i \le n} \in [0, 1]^d \times \mathbb{R}^p$ 
  - $\bullet$   $X_i \perp X_i$
  - $Y_i|(X_k)_k \perp Y_i|(X_k)_k$
  - Y|X has a density  $s_0$  w.r.t. Lebesgue measure
- **Regression** = specific modeling of the conditional density  $s_0(.|x)$
- Gaussian regression mixture with logistic weights:

$$s_{K,v,\Sigma,w}(y|x) = \sum_{k=1}^{K} \pi_{w,k}(x) \Phi_{v_k(x),\Sigma_k(x)}(y),$$

with 
$$\sigma_{w,k}(x) = \frac{e^{w_k(x)}}{\sum_{k'=1}^K e^{w_{k'}(x)}}$$
, logistic weights  $\Phi_{v_k(x), \sum_k(x)}$  density of  $\mathcal{N}(v_k(x), \sum_k(x))$ 

- Parameters:
  - K: number of components
  - v and  $\Sigma$ : K regression functions and covariance matrices functions
  - w: K weights functions defining the mixture proportions

#### Models

• Gaussian regression mixture with logistic weights:

$$s_{K,v,\Sigma,w}(y|x) = \sum_{k=1}^{K} \pi_{w,k}(x) \Phi_{v_k(x),\Sigma_k(x)}(y),$$

with 
$$\sigma_{w,k}(x) = \frac{e^{w_k(x)}}{\sum_{k'=1}^K e^{w_{k'}(x)}}$$
, logistic weights  $\sigma_{v_k(x),\Sigma_k(x)}$  density of  $\mathcal{N}(v_k(x),\Sigma_k(x))$ 

- Parameters:  $\theta = (K, v, \Sigma, w)$ 
  - K: number of components
  - v and  $\Sigma$ : K regression functions and covariance matrices functions
  - w: K weights functions defining the mixture proportions
- Model  $S_m = \{s_{\theta}, \theta \in \Theta_m\}$  with  $\Theta_m = \{K\} \otimes \Upsilon_K \otimes V_K \otimes W_k$ :
  - K: number of components.
  - $\Upsilon_K$  and  $V_K$ : sets for the K-tuple of regressions functions and covariance matrices functions.
  - $W_K$ : sets for for the K-tuple of weights functions.
- Typical choice:
  - $\Upsilon_K$  and  $W_K$ : tensorial product of polynomial sets of low degree.
  - $W_K$ : constant covariance structures independent of X.

# Maximum likelihood and penalization

- Model  $S_m = \{s_\theta, \theta \in \Theta_m\}$  with  $\Theta_m = \{K\} \otimes \Upsilon_K \otimes V_K \otimes W_K$ :
  - K: number of components.
  - $\Upsilon_K$  and  $V_K$ : sets for the K-tuple of regressions functions and covariance matrices functions.
  - $W_K$ : sets for for the K-tuple of weights functions.
- Maximum likelihood estimation within each model:

$$\widehat{s}_m = \underset{\theta \in \Theta_m}{\operatorname{argmax}} - \sum_{i=1}^n \ln s_{\theta}(Y_i|X_i)$$

• Model selection by a penalization proportional to the dimension:

$$\widehat{m} = \operatorname*{argmin}_{m \in \mathcal{M}} \sum_{k=1}^{K} - \ln \widehat{s}_m(Y_i | X_i) + \kappa \dim \Theta_m$$

Usual complexity/fidelity tradeoff.

#### Contributions

#### Characterization of the theoretical performances

- Penalty choice:  $pen(m) = \kappa(C + \ln n) \dim(S_m)$ .
- Oracle inequality:

$$\mathbb{E}\left[JKL_{\rho}^{\otimes n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq C_1 \inf_{m \in \mathcal{M}} \left(\inf_{s_m \in S_m} KL^{\otimes n}(s_0,s_m) + \frac{\operatorname{pen}(m)}{n}\right) + \frac{C_2}{n}$$

#### Numerical implementation of the penalized maximum likelihood

- EM type minimization scheme with a focus on initialization issues.
- Practical scheme for the penalty calibration with the slope heuristic approach.

# Conditional density and selection

- General framework: observation of  $(X_i, Y_i)$  with  $X_i$  independent and  $Y_i$  cond. independent of law of density  $s_0(y|X_i)$ .
- **Goal:** estimation of  $s_0(y|x)$ .
- Penalized model selection principle:
   choice of a collection of cond. dens. models S<sub>m</sub> = {s<sub>m</sub>(y|x)} with m ∈ S,
  - Maximum likelihood estimation of a cond. density  $\hat{s}_m$  for each model  $S_m$ :

$$\hat{s}_m = \underset{s_m \in S_m}{\operatorname{argmin}} - \sum_{i=1}^m \ln s_m(Y_i|X_i)$$

• Selection of a model 
$$\widehat{m}$$
 by 
$$\widehat{m} = \operatorname*{argmin}_{m \in \mathcal{S}} - \sum_{i=1}^n \ln \widehat{s}_m(Y_i|X_i) + \operatorname{pen}(m).$$

with pen(m) well chosen.

Typical oracle inequality result:

$$\mathbb{E}\left[d^2(s_0,\widehat{s}_{\widehat{m}})\right] \leq C\inf_{m \in \mathcal{S}}\left(\inf_{s_m \in \mathcal{S}_m} \mathsf{KL}(s_0,s_m) + \frac{\mathrm{pen}(m)}{n}\right) + \frac{C'}{n}.$$

Short bib.: Rosenblatt, Fan et al., de Gooijer and Zerom,
 Efromovitch, Brunel, Comte, Lacour... / Plugin, direct estimation,
 L<sup>2</sup>, minimax, censure...

# Ideal oracle inequality

Oracle inequality:

$$\mathbb{E}\left[\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq C_1 \inf_{S_m \in \mathcal{S}} \left( \underbrace{\inf_{\substack{s_m \in S_m \\ \text{Bias term}}} \mathit{KL}^{\otimes_n}(s_0,s_m)}_{\text{Variance term}} + \underbrace{\frac{\Delta(m)}{n}}_{\text{Variance term}} \right)$$

as soon as pen(m) is large enough

- Divergence adapted to the conditional density setting:
  - Divergence on the product density conditioned on the design (Kolaczyk, Bigot).
  - Tensorization principle and expectation on the design: design:

$$\mathit{KL} 
ightarrow \mathit{KL}^{\otimes_n}(s,s') = \mathbb{E}\left[ rac{1}{n} \sum_{i=1}^n \mathit{KL}\left(s(\cdot|X_i),s'(\cdot|X_i)
ight)
ight]$$

- Much more information using the second approach because losses used are larger.
- Ability to handle independent but non i.i.d. case and integrated loss.
- Classical density estimation theorem if  $s(\cdot|X_i) = s(\cdot)$ .

### **Notations**

- Let for any function g(x, y),
  - $P_n^{\otimes_n}(g)$ : its empirical process  $P_n^{\otimes_n}(g) = \frac{1}{n} \sum_{i=1}^n g(X_i, Y_i)$ .
  - $P^{\otimes_n}(g)$ : its expectation  $P^{\otimes_n}(g) = \mathbb{E}\left[P_n^{\otimes_n}(g)\right] = \mathbb{E}\left[\frac{1}{n}\sum_{i=1}^n g(X_i,Y_i)\right]$ .
  - $\nu_n^{\otimes_n}(g) = P_n^{\otimes_n}(g) P^{\otimes_n}(g)$ : its **recentered** process.
- Maximum likelihood estimate:

$$\widehat{s}_{m} = \underset{s_{m} \in S_{m}}{\operatorname{argmin}} \sum_{i=1}^{n} -\ln s_{m}(Y_{i}|X_{i}) = \underset{s_{m} \in S_{m}}{\operatorname{argmin}} P_{n}^{\otimes_{n}}(-\ln s_{m})$$
$$= \underset{s_{m} \in S_{m}}{\operatorname{argmin}} P_{n}^{\otimes_{n}}(-\ln \frac{s_{m}}{s_{0}})$$

Best projection:

$$\widetilde{s}_m = \underset{s_m \in S_m}{\operatorname{argmin}} KL^{\otimes_n}(s_0, s_m) = \underset{s_m \in S_m}{\operatorname{argmin}} P^{\otimes_n}(-\ln \frac{s_m}{s_0})$$

$$= \underset{s_m \in S_m}{\operatorname{argmin}} P^{\otimes_n}(-\ln s_m)$$

# Ideal penalty

By definition:

$$\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) = P_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_m}{s_0}\right) \underbrace{-\nu_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_m}{s_0}\right)}_{\mathrm{pen}_{\mathrm{id}}(m)/n}$$

• With the *ideal* penalty  $pen_{id}(m)$ :

$$\begin{split} \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) &= \mathit{P}_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) + \frac{\mathrm{pen}_{\mathrm{id}}(\widehat{m})}{n} \\ &\leq \inf_{m} \mathit{P}_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_m}{s_0} \right) + \frac{\mathrm{pen}_{\mathrm{id}}(m)}{n} \leq \inf_{m} \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) \\ &\leq \inf_{m} \left( \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m) + \left( \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) - \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m) \right) \end{split}$$

Ideal penalty oracle inequality:

$$\mathbb{E}\left[\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq \inf_{S_m \in \mathcal{S}} \left(\underbrace{\mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m)}_{\text{Bias term}} + \underbrace{\mathbb{E}\left[\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) - \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m)\right]}_{\text{Variance term}}\right)$$

# Non ideal penalization

By construction

$$KL^{\otimes_n}(s_0, \widehat{s}_{\widehat{m}}) = P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right)$$

$$= P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) + \frac{\operatorname{pen}(\widehat{m})}{n}$$

$$- \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

$$\leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_m}{s_0} \right) + \frac{\operatorname{pen}(m)}{n} \right)$$

$$- \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

# Non ideal penalization

• Using  $\widetilde{s}_m = \operatorname{argmin}_{s_m \in S_m} KL^{\otimes_n}(s_0, s_m)$ :

$$\begin{split} \mathit{KL}^{\otimes_n}(s,\widehat{s}_{\widehat{m}}) &\leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_m}{s_0} \right) + \frac{\mathrm{pen}(m)}{n} \right) \\ &- \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \frac{\mathrm{pen}(\widehat{m})}{n} \\ \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) &\leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{s_0} \right) + \frac{\mathrm{pen}(m)}{n} \right) \\ &- \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \frac{\mathrm{pen}(\widehat{m})}{n} \end{split}$$

# Non ideal penalization

Summary:

$$KL^{\otimes_n}(s_0, \widehat{s}_{\widehat{m}}) \leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{s_0} \right) + \frac{\operatorname{pen}(m)}{n} \right) - \nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

• Oracle inequality up to something:

$$\mathbb{E}\left[\mathsf{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq \min_{S_m \in \mathcal{S}} \left(\mathsf{KL}^{\otimes_n}(s_0,\widetilde{s}_m) + \frac{\mathrm{pen}(m)}{n}\right) \\ + \mathbb{E}\left[-\nu_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_{\widehat{m}}}{s_0}\right) - \frac{\mathrm{pen}(\widehat{m})}{n}\right]$$

- If  $\mathbb{E}\left[-\nu_n^{\otimes_n}\left(-\ln\frac{\widehat{s}_{\widehat{m}}}{s_0}\right)-\frac{\operatorname{pen}(\widehat{m})}{n}\right]\leq 0$  then **exact** oracle inequality!
- If  $\mathbb{E}\left[-\nu_n^{\otimes_n}\left(-\ln\frac{\widehat{s}_{\widehat{m}}}{s_0}\right) \frac{\operatorname{pen}(\widehat{m})}{n} \epsilon KL^{\otimes_n}(s_0, \widehat{s}_{\widehat{m}})\right] \leq 0$  then inexact oracle inequality.

### Kullback-Leibler and extension

• Issue in the previous approach: control of

$$\nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right)$$

hard due to the **unboundedness** of  $-\ln \frac{\widehat{s}_{m}}{s_{0}}$ • **Trick:** replace this quantity by the bounded one

$$-rac{1}{a}\lnrac{
ho\widehat{s}_{\widehat{m}}+(1-
ho)s_0}{s_0}\leq -rac{1}{a}\ln(1-
ho)$$

By convexity,

$$-rac{1}{
ho}\lnrac{
ho\widehat{s}_{\widehat{m}}+(1-
ho)s_0}{s_0}\leq -\lnrac{\widehat{s}_{\widehat{m}}}{s_0}$$

Jensen-Kullback-Leibler divergence:

$$\mathit{JKL}^{\otimes_n}_\rho(s_0,\widehat{s}_m) = P^{\otimes_n}\left(-\frac{1}{\rho}\ln\frac{\rho\widehat{s}_{\widehat{m}} + (1-\rho)s_0}{s_0}\right) = \frac{1}{\rho}\mathit{KL}^{\otimes_n}(s_0,\rho\widehat{s}_m + (1-\rho)s_0)$$

$$\leq P^{\otimes_n}\left(-\lnrac{\widehat{s}_{\widehat{m}}}{s_0}
ight) = \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m)$$

# $JKL^{\otimes_n}$ and non ideal penalization

By construction

$$JKL_{\rho}^{\otimes_{n}}(s_{0},\widehat{s}_{\widehat{m}}) = P_{n}^{\otimes_{n}} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}} \right)$$

$$- \nu_{n}^{\otimes_{n}} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}} \right)$$

$$\leq P_{n}^{\otimes_{n}} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_{0}} \right) - \nu_{n}^{\otimes_{n}} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}} \right)$$

$$\leq P_{n}^{\otimes_{n}} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_{0}} \right) + \frac{\operatorname{pen}(\widehat{m})}{n}$$

$$- \nu_{n}^{\otimes_{n}} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

$$\leq \min_{S_{m} \in \mathcal{S}} \left( P_{n}^{\otimes_{n}} \left( -\ln \frac{\widehat{s}_{m}}{s_{0}} \right) + \frac{\operatorname{pen}(m)}{n} \right)$$

$$- \nu_{n}^{\otimes_{n}} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

# $JKL^{\otimes_n}$ and non ideal penalization

• Using  $\widetilde{s}_m = \operatorname{argmin}_{s_m \in S_m} KL^{\otimes_n}(s_0, s_m)$ :

$$\begin{split} \textit{JKL}_{\rho}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) &\leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_m}{s_0} \right) + \frac{\mathrm{pen}(m)}{n} \right) \\ &- \nu_n^{\otimes_n} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_0}{s_0} \right) - \frac{\mathrm{pen}(\widehat{m})}{n} \\ \textit{JKL}_{\rho}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) &\leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{s_0} \right) + \frac{\mathrm{pen}(m)}{n} \right) \\ &- \nu_n^{\otimes_n} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_0}{s_0} \right) - \frac{\mathrm{pen}(\widehat{m})}{n} \end{split}$$

# $JKL^{\otimes_n}$ and non ideal penalization

Summary:

$$JKL_{\rho}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) \leq \min_{S_m \in \mathcal{S}} \left( P_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{s_0} \right) + \frac{\operatorname{pen}(m)}{n} \right) \\
- \nu_n^{\otimes_n} \left( -\frac{1}{\rho} \ln \frac{\rho \widehat{s}_{\widehat{m}} + (1-\rho)s_0}{s_0} \right) - \frac{\operatorname{pen}(\widehat{m})}{n}$$

Oracle inequality up to something:

$$\mathbb{E}\left[JKL_{\rho}^{\otimes_{n}}(s_{0},\widehat{s}_{\widehat{m}})\right] \leq \min_{S_{m}\in\mathcal{S}}\left(KL^{\otimes_{n}}(s_{0},\widetilde{s}_{m}) + \frac{\operatorname{pen}(m)}{n}\right) + \mathbb{E}\left[-\nu_{n}^{\otimes_{n}}\left(-\frac{1}{\rho}\ln\frac{\rho\widehat{s}_{\widehat{m}} + (1-\rho)s_{0}}{s_{0}}\right) - \frac{\operatorname{pen}(\widehat{m})}{n}\right]$$

 Under some assumptions on the model collection, it exists a penalty such that

$$\mathbb{E}\left[-\nu_n^{\otimes_n}\left(-\ln\frac{\widehat{s}_{\widehat{m}}}{s_0}\right) - \frac{\mathrm{pen}(\widehat{m})}{n} - \epsilon \textit{JKL}_{\rho}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq 0$$

• For such a penalty, one has an inexact oracle inequality.

#### Theorem

**Assumption (H)**: For every model  $S_m$  in the collection  $\mathcal{S}$ , there is a non-decreasing function  $\phi_m(\delta)$  such that  $\delta \mapsto \frac{1}{\delta}\phi_m(\delta)$  is non-increasing on  $(0,+\infty)$  and for every  $\sigma \in \mathbb{R}^+$  and every  $s_m \in S_m$ 

$$\int_0^\sigma \sqrt{H_{[\cdot],d^{\otimes_n}}(\epsilon,S_m(s_m,\sigma))}\,d\epsilon \leq \phi_m(\sigma).$$

**Assumption (K)**: There is a family  $(x_m)_{m \in \mathcal{M}}$  of non-negative number such that

$$\sum_{m\in\mathcal{M}}e^{-x_m}\leq \Sigma<+\infty$$

#### **Theorem**

Assume we observe  $(X_i, Y_i)$  with unknown conditional  $s_0$ . Let  $\mathcal{S} = (S_m)_{m \in \mathcal{M}}$  a at most countable collection of conditional density sets. Assume Assumptions (H), (K) and (S) hold.

Let  $\hat{s}_m$  be a  $\delta$  -log-likelihood minimizer in  $S_m$ :

$$\sum_{i=1}^{n} -\ln(\widehat{s}_m(Y_i|X_i)) \leq \inf_{s_m \in S_m} \left(\sum_{i=1}^{n} -\ln(s_m(Y_i|X_i))\right) + \delta$$

Then for any  $\rho \in (0,1)$  and any  $C_1 > 1$ , there is a constant  $\kappa_0$  depending only on  $\rho$  and  $C_1$  such that, as soon as for every index  $m \in M$  pen $(m) > \kappa(\mathfrak{D}_m + \chi_m)$  with  $\kappa > \kappa_0$ 

as soon as for every index  $m \in \mathcal{M}$   $\operatorname{pen}(m) \ge \kappa(\mathfrak{D}_m + x_m)$  with  $\kappa > \kappa_0$  where  $\mathfrak{D}_m = n\sigma_m^2$  with  $\sigma_m$  the unique root of  $\frac{1}{\sigma}\phi_m(\sigma) = \sqrt{n}\sigma$ ,

the penalized likelihood estimate  $\widehat{s}_{\widehat{m}}$  with  $\widehat{m}$  defined by

$$\widehat{m} = \underset{m \in \mathcal{M}}{\operatorname{argmin}} \sum_{i=1}^{n} - \ln(\widehat{s}_{m}(Y_{i}|X_{i})) + \operatorname{pen}(m)$$

$$\textit{satisfies} \qquad \mathbb{E}\left[\textit{JKL}_{\rho}^{\otimes_n}(s_0,\widehat{s_m})\right] \leq C_1\left(\inf_{S_m \in S}\left(\inf_{s_m \in S_m}\textit{KL}^{\otimes_n}(s_0,s_m) + \frac{\mathrm{pen}(m)}{n}\right) + \frac{\kappa_0\Sigma + \delta}{n}\right).$$

# Simplified Theorem...

#### Oracle inequality:

$$\mathbb{E}\left[JKL_{\rho}^{\otimes_{n}}(s_{0},\widehat{s}_{\widehat{m}})\right] \leq C_{1}\left(\inf_{S_{m}\in\mathcal{S}}\left(\inf_{s_{m}\in\mathcal{S}_{m}}KL^{\otimes_{n}}(s_{0},s_{m}) + \frac{\mathrm{pen}\;m}{n}\right) + \frac{\kappa_{0}\Sigma + \delta}{n}\right)$$

as soon as

$$pen(m) \ge \kappa (\mathfrak{D}_m + x_m)$$
 with  $\kappa > \kappa_0$ ,

where  $\mathfrak{D}_m$  measure the **complexity of the model**  $S_m$  (entropy term) and  $x_m$  the **coding cost** within the collection.

- $\mathfrak{D}_m$  linked to the **bracketing entropy** of  $S_m$  with respect to the tensorized Hellinger distance  $d^{2\otimes n}$ .
- Often  $\mathfrak{D}_m \propto (\log n) \dim(S_m)...$

# Penalty and complexities

Control required on

$$-\nu_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_{\widehat{m}}}{s_0}\right) - \frac{\operatorname{pen}(\widehat{m})}{n} - \epsilon JKL_\rho^{\otimes_n}(s_0, \widehat{s}_{\widehat{m}})$$

through a **supremum**!

- Control in expectation requires a pen(m) taking into account
  - the intrinsic complexity of the model,
  - the complexity of the collection.
- Here:
  - Model complexity: entropy complexity  $\mathfrak{D}_m$  defined from the *bracketing* entropy  $H_{[\cdot],d^{\otimes_n}}(\epsilon,S_m)$  of  $S_m$  with respect to the tensorized Hellinger distance  $d^{2\otimes_n}$ .
  - Collection (coding): Kraft type inequality  $\sum_{m \in S} e^{-x_m} \leq \Sigma < +\infty$
- Classical constraint on the penalty

$$pen(m) \ge \kappa (\mathfrak{D}_m + x_m)$$
 with  $\kappa > \kappa_0$ .

• Often  $\mathfrak{D}_m \propto (\ln(n)) \dim(S_m)$  and thus classical penalization by dimension setting...

# Brackets and complexity

- Bracketing entropy:  $H_{[\cdot],d^{\otimes_n}}(\epsilon,S) = \text{logarithm of the minimum}$  number of brackets  $[t_i^-,t_i^+]$  such that
  - $\forall i, d^{\otimes_n}(t_i^-, t_i^+) \le \epsilon$   $\forall s \in S, \exists i, t_i^- < s < t_i^+$ 
    - where  $d^{\otimes_n} = \sqrt{d^{2\otimes_n}} = \sqrt{\mathbb{E}\left[\frac{1}{n}\sum d^2(s(\cdot|X_i),s'(\cdot|X_i))\right]}$  is the tensorized Hellinger distance.
- **Assumption** (H): for all model  $S_m$ , there is a non decreasing  $\phi_m(\delta)$  such that  $\delta \mapsto \frac{1}{\delta}\phi_m(\delta)$  is non increasing  $(0, +\infty)$  and such that for all  $\sigma \in \mathbb{R}^+$  and all  $s_m \in S_m$

$$\int_0^\sigma \sqrt{H_{[\cdot],d^{\otimes n}}(\epsilon,S_m)}\,d\epsilon \leq \phi_m(\sigma),$$

- Complexity  $\mathfrak{D}_m$  def. as  $n\sigma_m^2$  with  $\sigma_m$  unique root of  $\phi_m(\sigma) = \sqrt{n}\sigma^2$ .
- Key: Dudley type integral and optimization of a deviation bound.
- Typically,  $H_{[\cdot],d^{\otimes n}}(\epsilon,S_m)$   $\sim \dim S_m(C+\log 1/\epsilon)$  which implies  $\mathfrak{D}_m \propto (\ln n) \dim(S_m)...$

### Gaussian regression mixtures

- Model  $S_m = \{s_\theta, \theta \in \Theta_m\}$  with  $\Theta_m = \{K\} \otimes \Upsilon_K \otimes V_K \otimes W_k$ :
  - $\Upsilon_K$  and  $V_K$ : sets for the K-tuple of regressions functions and covariance matrices functions.
  - $W_K$ : sets for for the K-tuple of weights functions.
- Structural assumptions:
  - $V_K$  is a set of covariance matrices independent of the covariate,
  - $\bullet$   $\Upsilon_K$  and  $W_K$  are such that

$$\begin{split} & H_{\max_{k=1}^K \sup_x \|\cdot_k(x)\|}(\delta, W_K) \leq \dim(W_K) \left( C_W + \ln \frac{1}{\delta} \right) \\ & H_{\max_{k=1}^K \sup_x \|\cdot_k(x)\|_2}(\delta, \Upsilon_K) \leq \dim(\Upsilon_K) \left( C_\Upsilon + \ln \frac{1}{\delta} \right) \end{split}$$

- Satisfied for instance if  $\Upsilon_K$  and  $W_K$  are K-tuples of **polynomials** with bounded coefficients and x is bounded.
- Th: Under this assumption, if  $pen(m) = \kappa(C + \ln n) \dim(S_m)$  then

$$\mathbb{E}\left[JKL_{\rho}^{\otimes n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq C_1 \inf_{m \in \mathcal{M}} \left(\inf_{s_m \in S_m} KL^{\otimes n}(s_0,s_m) + \frac{\operatorname{pen}(m)}{n}\right) + \frac{C_2}{n}$$

• **Key:** upper bound of the bracketing entropy  $H_{[\cdot],d^{\otimes_n}}(\epsilon,S_m)$ .

# Bracketing entropy decomposition

Model:

$$S_{m} = \{ \sum_{k=1}^{K} \pi_{w,k}(x) \Phi_{v_{k}(x), \Sigma_{k}(x)}(y), (K, v, \Sigma, w) \in \Theta_{m} \}$$

with  $\Theta_m = \{ K \} \otimes \Upsilon_K \otimes V_K \otimes W_k$ 

Weight and regression models:

$$\mathcal{W}_{K} = \left\{ \left( \pi_{w,k}(x) \right)_{k=1}^{K}, w \in W_{K} \right\}$$

$$\mathcal{R}_{K} = \left\{ \left( \Phi_{v_{k}(x), \Sigma_{k}(x)}(y) \right)_{k=1}^{K}, (v, \Sigma) \in \Upsilon_{K} \times V_{K} \right\}$$

Splitting properties:

$$H_{[.],d^{\otimes_n}}(\delta,S_m) \leq H_{[.],\sup_{\kappa}\max_{\kappa}d}\left(\frac{\delta}{5},\mathcal{R}_{\mathcal{K}}\right) + H_{[.],\sup_{\kappa}d}\left(\frac{\delta}{5},\mathcal{W}_{\mathcal{K}}\right)$$

# Bracketing entropy decomposition

• Gaussian *K*-tuple bracketing entropy:

$$\begin{split} H_{[.],\sup_{x}\max_{k}d}\left(\frac{\delta}{5},\mathcal{R}_{\mathcal{K}}\right) &\leq H_{\max_{k=1}^{\mathcal{K}}\sup_{x}\|\cdot_{k}(x)\|_{2}}(\epsilon_{1}\delta,\Upsilon_{\mathcal{K}}) + H_{d}(\epsilon_{2}\delta,V_{\mathcal{K}}) \\ &\leq \left(\dim(\Upsilon_{\mathcal{K}}) + \dim(V_{\mathcal{K}})\right)\left(C_{\mathcal{R}} + \ln\frac{1}{\delta}\right) \end{split}$$

• Logistic weight *K*-tuple bracketing entropy:

$$H_{[.],\sup_{\kappa} d}\left(\frac{\delta}{5}, \mathcal{W}_{K}\right) \leq H_{\max_{k}\sup_{\kappa}\|\cdot_{k}\|}\left(\frac{\epsilon_{3}\delta}{\sqrt{K}}, \mathcal{W}_{K}\right) \leq \dim(\mathcal{W}_{K})\left(C_{W} + \ln\frac{1}{\delta}\right)$$

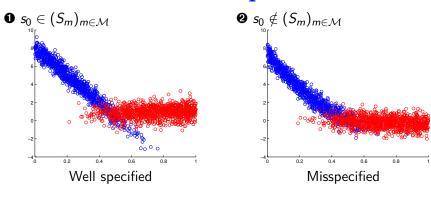
• Bracketing entropy bound if  $K \leq K_{max}$ :

$$H_{[.],d^{\otimes n}}(\delta, S_m) \leq H_{[.],\sup_{\kappa}\max_{k} d}\left(\frac{\delta}{5}, \mathcal{R}_{K}\right) + H_{[.],\sup_{\kappa} d}\left(\frac{\delta}{5}, \mathcal{W}_{K}\right)$$

$$\leq \left(\dim(\Upsilon_{K}) + \dim(V_{K}) + \dim(W_{K})\right)\left(C + \ln\frac{1}{\delta}\right)$$

$$\leq \dim(S_m)\left(C + \ln\frac{1}{\delta}\right)$$

# Numerical experiments



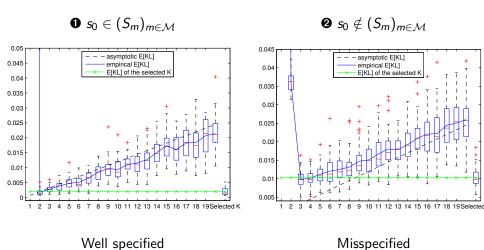
2 000 points

- Models  $S_m$  used:
  - Affine models for the weights and the regressions:

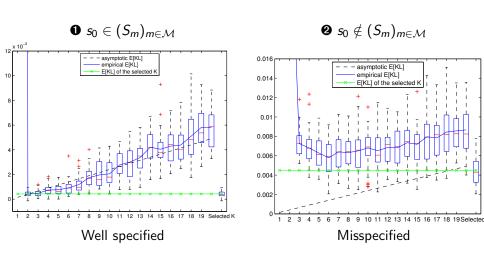
$$\Upsilon_{K} = W_{K} = \{(a_{k}x + b_{k})_{k=1}^{K}, (a, b) \in \mathbb{R}^{K \times 2}\}$$

- Free variance:  $V_K = \mathbb{R}_+^K$
- Only choice is the number of components K

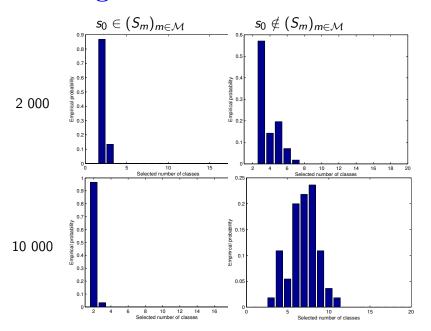
# KL risk 2 000 points



# KL risk 10 000 points



# Histograms of the selected K



# Numerical optimization

- Model  $S_m = \{s_\theta, \theta \in \Theta_m\}$  with  $\Theta_m = \{K\} \otimes \Upsilon_K \otimes V_K \otimes W_k$ :
  - *K*: number of components.
  - $\Upsilon_K$  and  $V_K$ : sets for the K-tuple of regressions functions and covariance matrices functions.
  - $W_K$ : sets for the K-tuple of weights functions.
- Maximum likelihood estimation:

$$\widehat{s}_m = \underset{\theta \in \Theta_m}{\operatorname{argmin}} - \sum_{i=1}^N \ln s_{\theta}(Y_i|X_i)$$

Penalized model selection:

$$\widehat{m} = \underset{m}{\operatorname{argmin}} - \sum_{i=1}^{N} \ln \widehat{s}_{m}(Y_{i}|X_{i}) + \kappa \operatorname{dim} \Theta_{m}$$

- Model selection computed by exhaustive exploration.
- Focus on maximum likelihood estimation!

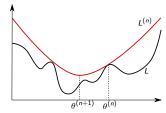
### Maximum likelihood estimation

- Model  $S_m = \{s_\theta, \theta \in \Theta_m\}$  with  $\Theta_m = \{K\} \otimes \Upsilon_K \otimes V_K \otimes W_k$ :
  - *K*: number of components.
  - $\Upsilon_K$  and  $V_K$ : sets for the K-tuple of regressions functions and covariance matrices functions.
  - $W_K$ : sets for the K-tuple of weights functions.
- Maximum likelihood estimation:

$$\widehat{s}_{m} = \underset{\theta \in \Theta_{m}}{\operatorname{argmin}} \underbrace{-\sum_{i=1}^{N} \operatorname{In} \left( \sum_{k=1}^{K} \pi_{w,k}(X_{i}) \Phi_{v_{k}(X_{i}), \Sigma_{k}(X_{i})}(Y_{i}) \right)}_{L(\theta)}$$

- Non convex minimization problem!
- Majorization/Minimization approach

# MM approach



- Iterative approach to minimize  $L(\theta)$  by minimizing a sequence of (convex) proxies of L.
- Majorization/Minimization:
  - **Current estimate** of the minimizer:  $\theta^{(n)}$
  - Construction of a **Majorization**  $L^{(n)}$  of L such that  $L^{(n)}(\theta^{(n)}) = L(\theta^{(n)})$  with  $L^{(n)}$  easy to minimize (convex for example).
  - Computation of a Minimizer

$$\theta^{(n+1)} = \operatorname{argmin} L^{(n)}(\theta)$$

- By construction,  $L(\theta^{(n+1)}) \leq L(\theta^{(n)})!$
- Very generic methodology...
- Minimization can be replaced by a diminution...

### Maximum Likelihood and EM

Back to our maximum likelihood:

$$L(\theta) = L(K, v, \mathbf{\Sigma}, w) = -\sum_{i=1}^{N} \ln \left( \sum_{k=1}^{K} \pi_{w,k}(X_i) \Phi_{v_k(X_i), \mathbf{\Sigma}_k(X_i)}(Y_i) \right)$$

• **EM**: specific case of MM for this type of mixture.

• (Conditional) Expectation: at step 
$$n$$
, we let

$$P_k^{i,(n)} = P\left(k_i = k \middle| X_i, Y_i, \frac{K}{K}, \upsilon^{(n)}, \Sigma^{(n)}, \frac{W^{(n)}}{M}\right)$$

and 
$$L^{(n)}(K, v, \Sigma, \mathbf{w}) = -\sum_{k=1}^{N} \sum_{k=1}^{K} P_{k}^{i,(n)} \ln \left( \pi_{\mathbf{w},k}(X_{i}) \Phi_{v_{k}(X_{i}), \Sigma_{k}(X_{i})}(Y_{i}) \right).$$

• Maj. prop.:  $L < L^{(n)} + Cst^{(n)}$  with equ. at

$$\theta = (K, v^{(n)}, \Sigma^{(n)}, w^{(n)})$$
• Separability in  $(v^{(n)}, \Sigma^{(n)})$  and  $w^{(n)}$ :

$$L^{(n)}(K, v, \Sigma, w) = \left(-\sum_{i=1}^{N} \sum_{k=1}^{K} P_k^{i,(n)} \ln \Phi_{v_k(X_i), \Sigma_k(X_i)}(Y_i)\right) + \left(-\sum_{i=1}^{N} \sum_{k=1}^{K} P_k^{i,(n)} \ln \pi_{w,k}(X_i)\right)$$

# Minimization of $L^{(n)}$

• Separability in  $(v^{(n)}, \Sigma^{(n)})$  and  $w^{(n)}$ :

$$L^{(n)}(\mathbf{K}, v, \mathbf{\Sigma}, \mathbf{w}) = \left(-\sum_{i=1}^{N} \sum_{k=1}^{K} P_k^{i,(n)} \ln \Phi_{v_k(X_i), \mathbf{\Sigma}_k(X_i)}(Y_i)\right) + \left(-\sum_{i=1}^{N} \sum_{k=1}^{K} P_k^{i,(n)} \ln \pi_{\mathbf{w}, k}(X_i)\right)$$

- For the regression parameters  $(v^{(n)}, \Sigma^{(n)})$ :
  - K weighted linear regressions:  $-\sum_{i=1}^{N} P_k^{i,(n)} \ln \Phi_{\upsilon_k(X_i),\Sigma_k(X_i)}(Y_i)$
  - Explicit formulas!
- For the weight parameters  $w^{(n)}$ :
  - Single K modality logistic regression:  $-\sum_{k=1}^{N}\sum_{k=1}^{K}P_{k}^{i,(n)}\ln\pi_{w,k}(X_{i})$
  - Iterative minimization scheme (Newton = Iterative Reweighted Least Square)

#### Initialization

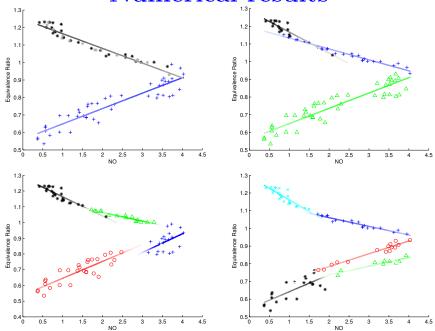
- Very important issue!
- For the **weights:** initialization to uniform weights seems sufficient.
- For the means:
  - Comparison between several strategies
    - Naive purely random initialization
    - Small-EM: Random initialization followed by a few minimization steps and selection
    - Advanced Small-EM: Initialization based on a first 2D clustering followed by a few minimization steps and selection
    - Advanced Small-EM 2: Initialization based on a random drawing of lines between points clustering followed by a few minimization steps and selection
  - Criterion: lowest likelihood for a given amount of time!
  - Similar results in term on expectation but different behaviors in term of dispersion:
    - Too simple strategies fail sometimes to provide a satisfactory answer while too complex ones may not explore sufficient local maxima.
    - Winner: Advanced Small-EM 2 with 3 minimizations steps and 50 candidates

#### Newton-EM Algorithm

#### Initialization with Advanced Small-EM 2:

- Initialization based on a random drawing of lines between points clustering
- 3 minimizations steps
- selection among 50 candidates
- Iterate until convergence:
  - Newton steps over weights w<sup>(n)</sup> if the likelihood increases (up to 5 times)
  - K linear regressions to update mean and variance  $(v^{(n)}, \Sigma^{(n)})$  in each class
- Note: Initialization issues in high dimension with this scheme!

# Numerical results



### Penalization strategy

Penalized model selection:

$$\widehat{m} = \operatorname*{argmin}_{m \in \mathcal{M}} \sum_{k=1}^{K} - \ln \widehat{s}_m(Y_i|X_i) + \operatorname{pen}(m)$$

Theoretical analysis:

$$pen(m) = \kappa(C + \ln n) \dim \Theta_m$$

- $\kappa$  and C are only loosely upper bounded!
- In practice, use  $pen(m) = \kappa \dim \Theta_m$  with  $\kappa$  chosen appropriately.
- Classical choice:
  - BIC:  $\kappa = \log n/2$
  - AIC:  $\kappa = 1$
- Here: **Jump/slope heuristic** = data driven choice of  $\kappa$

#### Ideal penalty

By definition:

$$\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) = P_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_m}{s_0}\right) \underbrace{-\nu_n^{\otimes_n} \left(-\ln\frac{\widehat{s}_m}{s_0}\right)}_{\mathrm{pen}_{\mathrm{id}}(m)/n}$$

• With the *ideal* penalty pen<sub>id</sub>(m):

$$\begin{split} \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}}) &= P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) + \frac{\mathrm{pen}_{\mathrm{id}}(\widehat{m})}{n} \\ &\leq \inf_m P_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_m}{s_0} \right) + \frac{\mathrm{pen}_{\mathrm{id}}(m)}{n} \leq \inf_m \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) \\ &\leq \inf_m \left( \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m) + \left( \mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) - \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m) \right) \end{split}$$

Ideal penalty oracle inequality:

$$\mathbb{E}\left[\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_{\widehat{m}})\right] \leq \inf_{S_m \in \mathcal{S}} \left(\underbrace{\mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m)}_{\text{Bias term}} + \underbrace{\mathbb{E}\left[\mathit{KL}^{\otimes_n}(s_0,\widehat{s}_m) - \mathit{KL}^{\otimes_n}(s_0,\widetilde{s}_m)\right]}_{\text{Variance term}}\right)$$

# Jump/Slope heuristic

• Ideal penalty decomposition:

$$\begin{split} \frac{\mathrm{pen}_{\mathrm{id}}(m)}{n} &= -\nu_n^{\otimes_n} \left( -\ln \frac{\widehat{s}_{\widehat{m}}}{s_0} \right) \\ &= \nu_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{\widehat{s}_m} \right) - \nu_n^{\otimes_n} \left( -\ln \widetilde{s}_m \right) + \underbrace{\nu_n^{\otimes_n} (-\log s_0)}_{\text{independent of } m}. \end{split}$$

- Jump/Slope heuristic:
  - Concentration:  $\nu_n^{\otimes_n} \left( -\ln \widetilde{s}_m \right) \ll \nu_n^{\otimes_n} \left( -\ln \widetilde{\widetilde{s}_m} \right)$
  - Symmetry:  $P_n^{\otimes_n} \left( -\ln(\widetilde{s}_m/\widehat{s}_m) \right) \sim P^{\otimes_n} \left( -\ln(\widetilde{s}_m/\widetilde{s}_m) \right)$
- Resulting approximation:

$$\frac{\mathrm{pen_{id}}(m)}{n} \sim 2P_n^{\otimes_n} \left(-\ln\frac{\widetilde{s}_m}{\widehat{s}_m}\right) \underbrace{-P^{\otimes_n}(-\log s_0)}_{\text{independent of }m}.$$

•  $P_n^{\otimes_n} \left( - \ln \frac{\widetilde{s}_m}{\widetilde{s}_m} \right)$  has still **to be estimated!** 

#### Minimal penalty

• If  $pen(m) = \kappa P_n^{\otimes_n} \left(-\ln \frac{\widetilde{s}_m}{\widehat{s}_m}\right)$  then

$$P_n^{\otimes_n}(-\ln\widehat{s}_m) + \operatorname{pen}(m) = (1-\kappa)P_n^{\otimes_n}(-\ln\widehat{s}_m) + \kappa P_n^{\otimes_n}(-\ln\widetilde{s}_m)$$

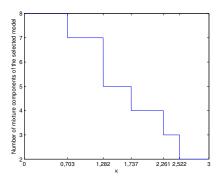
- No tradeoff is  $\kappa < 1!$
- Minimal penalty:  $pen_{\min}(m) = P_n^{\otimes_n} (-\ln(\widetilde{s}_m/\widehat{s}_m))$
- $\bullet$  Jump/Slope heuristic strongest assumption: parametric approximation of  $\mathrm{pen}_{\min}$

$$pen_{\min}(m) = pen(\kappa, m)$$

where pen shape is given by the theoretical study!

• Simplest case:  $pen(\kappa, m) = \kappa \dim S_m$ .

### Jump heuristic

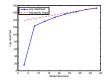


• Minimal penalty for which there is a **tradeoff**:

$$pen_{\min}(m) = pen(\kappa, m)$$

- Compute the models selected for several  $\kappa$  and detect **a jump in** the model dimensions.
- Not always a clear single jump...

## Slope heuristic



Observation:

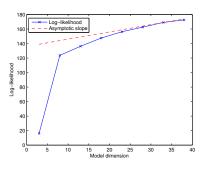
$$\operatorname{pen}_{\operatorname{id}}(m) = P_n^{\otimes_n} \left( -\ln \frac{\widetilde{s}_m}{\widehat{s}_m} \right) = P_n^{\otimes_n} \left( \ln \widehat{s}_m \right) + P_n^{\otimes_n} \left( -\ln \widetilde{s}_m \right)$$

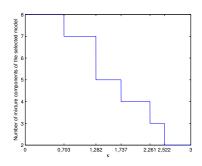
- If the model are more and more complex, one may expect that the **projection** bias converges to a constant:  $P_n^{\otimes_n}(-\ln \tilde{s}_m) \sim C$
- This implies  $\operatorname{pen}_{\operatorname{id}}(m) \sim P_n^{\otimes_n}(\ln \widehat{s}_m) + C$
- If  $\mathrm{pen}_{\mathrm{id}}(\mathit{indm}) = \mathrm{pen}(\kappa, m)$  then  $\kappa$  can be estimated by a regression as

$$pen(\kappa, m) - C \sim \underbrace{P_n^{\otimes_n}(\ln \widehat{s}_m)}_{\text{data driven}}$$

• If  $pen(\kappa, m) = \kappa \dim S_m$ ,  $\kappa$  measures the **slope** of  $P_n^{\otimes_n}(\ln \widehat{s}_m)$  with respect to dim  $S_m$ .

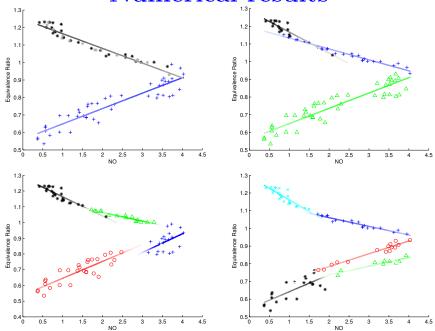
#### Slope heuristic



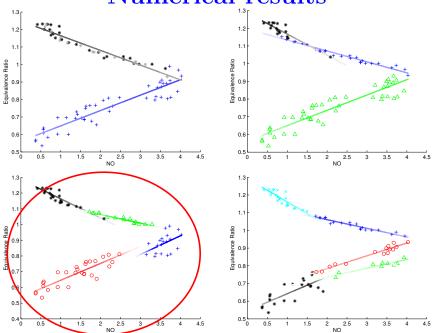


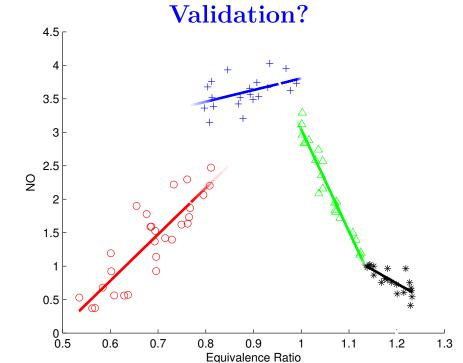
- Slope heuristic with pen $(\kappa, m) = \kappa \dim S_m$ :  $\kappa \sim 1$
- Resulting penalties:
  - Slope heuristic:  $pen(m) = 2 dim(S_m)$
  - BIC:  $pen(m) = 2,23 dim(S_m)$
  - AIC:  $pen(m) = 2 dim(S_m)$
- Selected number of clusters:
  - Slope heuristic: 4
  - BIC: 4
  - AIC: 7!

# Numerical results



# Numerical results





#### Conclusion

#### Framework:

- Mixture of regressions.
- Proposed tool: Mixture of Gaussian regressions with logistic mixing weights.
- Penalized maximum likelihood conditional density estimation.

#### Contributions:

- Theoretical guarantee for the conditional density estimation problem.
- Efficient minimization algorithm.
- Numerical penalty calibration.

#### Perspectives:

- Proof for penalty calibration by slope heuristic.
- Enhanced Spatialized Gaussian Mixture Model with piecewise logistic weights (S. Cohen).